

## ORIGINAL ARTICLE

## Updated gestational age specific birth weight, crown-heel length, and head circumference of Chinese newborns

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*Arch Dis Child Fetal Neonatal Ed* 2003;**88**:F229–F236**Objective:** To construct gestation specific standards of birth weight, crown-heel length, and head circumference of Chinese infants.**Design:** A prospective cross sectional population study.**Methods:** The birth weight, crown-heel length, and head circumference were prospectively measured using standard equipment in newborns delivered at 24–42 weeks gestation in the maternity units of 10 public hospitals and two private hospitals in Hong Kong. The findings were used to construct gestation specific standards of these variables. The LMS method using maximum penalised likelihood was used to perform model fitting. The results were compared with those obtained from a cohort of infants born in the same locality between 1982 and 1986.**Results:** From October 1998 to September 2000, a total of 10 032 infants were measured, representing 9.6% of the total deliveries in Hong Kong during that period. An extra 307 infants with gestation  $\leq 35$  weeks were recruited from October 2000 to June 2001. Each of the three variables showed a normal distribution at each gestational week. Gestation specific reference standards for each variable were constructed for male and female infants separately. Comparison with the 1982–1986 cohort showed a significant secular trend to increased birth weight. The trend was small, but significant, for crown-heel length and head circumference.**Conclusion:** These growth standards will provide useful references for the care of newborns of ethnic Chinese origin. These standards, especially that for birth weight, should be updated regularly.

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**B**irth weight, head circumference, and body length of newborn infants are important clinical indicators widely used for evaluation of prenatal growth and identification of infants that require detailed assessment and close monitoring during the neonatal period. Infants whose birth weights are too low or too high have higher mortality and morbidity than those of appropriate weight for gestation and increased risk of complications such as peripartum asphyxia, birth trauma, congenital malformations, and hypoglycaemia.<sup>1–5</sup> Body length is also of prognostic significance: an infant who is underweight but of normal body length has normal growth potential, but a small infant with small body length probably has impaired growth potential because of genetic factors or infectious or other teratogenic insults in early fetal life.<sup>6,7</sup> A recent study has shown that body length is also a predictor of perinatal mortality, with long infants being at higher risk of perinatal death.<sup>8,9</sup> Infants born with excessively small or large heads may have malformation of the central nervous system secondary to genetic or chromosomal abnormalities or teratogenic insults that carry grave prognostic implications.

Hong Kong, a British colony until July 1997, is a special administrative region of The People's Republic of China. Situated on the southern coast of China, 95% of its 6.8 million population are ethnic Chinese, most of whom are descendants of migrants from the southern provinces of China.<sup>10</sup> The reference standards for birth weight, crown-heel length, and head circumference used in the region were those provided by Lubchenco *et al*<sup>11,12</sup> until 1987, when local reference charts became available. These local references were based on measurements obtained from 8445 ethnic Chinese newborns born in Hong Kong between November 1982 and January 1986.<sup>13</sup> Over the past decade, the region has evolved from an industrial city to a commercial and financial centre. Associ-

ated with this change has been a rapid improvement in the standard of living and health indices, including infant and neonatal mortality. At the same time, the population has experienced an increasing growth rate from 1.1% in 1991 to 2.41% in 1996, and a falling birth rate from 11.7 per thousand in 1991 to 7.4 per thousand in 1999. Most of the population increase has been due to net inflow of people, mainly from Southern China, which accounted for 24.7–87.9% of the net population increase in the 1990s.<sup>14</sup> Over this time, health workers caring for newborns have noticed a trend of increasing size at birth of the local infants. A pilot study that we performed on 1350 full term newborns delivered in the Prince of Wales Hospital from January 1996 to April 1997 showed that the infants had greater gestation specific mean birth weight, crown-heel length, and head circumference than 15 years previously. In the light of these observations, a working group was formed to carry out a prospective study in 12 maternity units in the territory to establish a set of updated references for local infants.

**SUBJECTS AND METHODS**

The study was approved by the ethics committee on clinical research of The Chinese University of Hong Kong, and the ethics committees of the participating hospitals. In Hong Kong, all births take place in the maternity units of 20 hospitals. During the study period, about 70% of the newborns were delivered in the 10 public hospitals; the remaining 30% were born in the 10 private hospitals. To ensure that the sample selected truly represented the newborn population in Hong Kong, the babies were recruited from the maternity units of all 10 public hospitals and two randomly selected private hospitals.

**Table 1** Characteristics of the parents of the study infants

	Mother	Father
Education		
No formal education	114 (1.1)	31 (0.3)
Primary school	1179 (11.4)	941 (9.1)
Secondary school	7930 (76.7)	7403 (71.6)
University	1106 (10.7)	1964 (19)
Occupation		
Unemployed (including housewife)	6152 (59.5)	186 (1.8)
Professional	352 (3.4)	662 (6.4)
Non-manual worker	3256 (31.5)	3453 (33.4)
Manual worker	579 (5.6)	6038 (58.4)
Smoking		
Non-smoker	9584 (92.7)	5945 (57.5)
Smoker	755 (7.3)	4394 (42.5)
• Quitted during pregnancy	538 (5.2)	
• Continued to smoke during pregnancy	217 (2.1)	
Alcohol consumption		
Never	10019 (96.9)	7434 (71.9)
Social drinking	248 (2.4)	2429 (23.5)
Regular	72 (0.7)	476 (0.7)

Values in parentheses are percentages.

**Table 2** Birth weight (g), crown-heel length (cm), and head circumference (cm) for boys

Gestation (weeks)	n	Birth weight (g)		Crown-heel length (cm)		Head circumference (cm)	
		Mean	SD	Mean	SD	Mean	SD
24	9	664	87	322	31	226	5
25	16	775	87	345	18	237	8
26	12	886	64	352	13	248	10
27	19	1109	184	373	14	253	15
28	27	1156	168	378	20	262	14
29	26	1303	200	403	25	271	16
30	51	1476	190	408	18	281	14
31	41	1640	261	427	22	290	15
32	36	1896	399	434	28	300	20
33	89	2057	374	447	27	307	16
34	101	2234	354	455	21	312	13
35	148	2514	415	472	21	321	15
36	275	2803	419	482	20	331	14
37	432	3053	413	493	19	336	11
38	1054	3204	402	500	17	341	12
39	1304	3291	383	505	17	343	11
40	1179	3415	400	511	17	347	12
41	544	3518	429	514	17	350	12
42	115	3520	416	516	18	349	12

All measurements were carried out by two teams of field workers, each consisting of two investigators who had received training in the use of all measuring equipment. The precision of their measurements was assessed by establishing the interobserver agreement of the measurements obtained from the first 100 infants. In random sequence, the teams were stationed in each of the participating public hospitals for two months, and attempts were made to capture all eligible infants born during that period. Thus the study would capture about one sixth of the annual deliveries in each of the hospitals. The antenatal history and condition of each infant were carefully evaluated. A data sheet was used to document parental data, as well as the medical and pregnancy history of the mothers. To obtain a reasonable sample of infants born in the private hospitals, measurements in the two participating private hospitals took place over one year. Logistically it was not possible to include more private hospitals, in which the newborn infants were under the care of a large number of private obstetricians and paediatricians.

The main study lasted for two years from October 1998 to September 2000. At the end of two years, it was realised that

the number of infants < 35 weeks gestation was relatively small. The study was then extended for nine months until June 2001 to enroll more preterm infants.

#### Inclusion and exclusion criteria

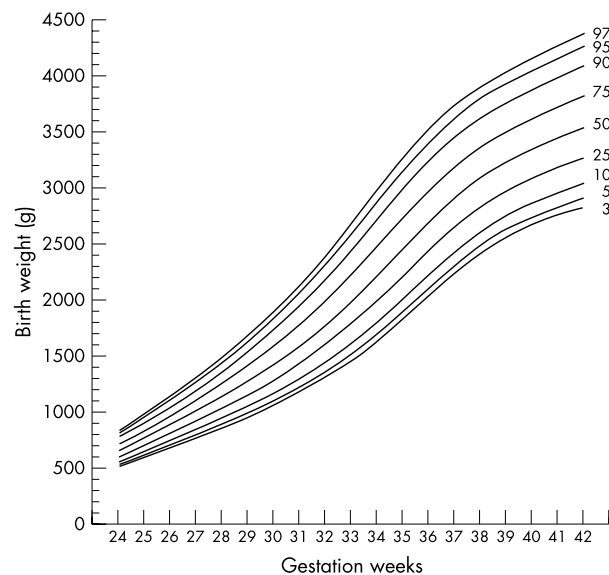
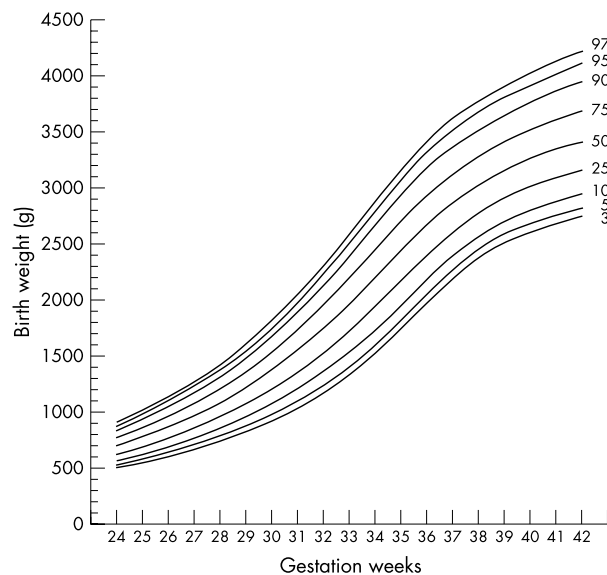
Singleton newborns of ethnic Chinese origin of gestation 24–43 weeks were eligible for the study provided that informed consent was given by the parents. Infants with the following conditions were excluded: moribund condition at birth; major congenital malformations; chromosomal abnormalities; gestational age impossible to determine. Infants born to mothers with medical conditions or complications of pregnancy were not excluded because the aim of the study was to construct community at large centile charts rather than those of a “healthy” population.

#### Assessment of gestational age

Gestational age was calculated in completed weeks from the findings of an early dating ultrasound performed before 20 weeks gestation. When this was not available, it was calculated from the last menstrual date if the mother had

**Table 3** Birth weight (g), crown-heel length (cm) and head circumference (cm) for girls

Gestation (weeks)	n	Birth weight (g)		Crown-heel length (cm)		Head circumference (cm)	
		Mean	SD	Mean	SD	Mean	SD
24	12	734	44	333	9	234	17
25	16	786	146	354	11	241	16
26	11	803	101	355	13	238	8
27	18	935	174	358	17	250	8
28	19	1116	170	376	22	257	12
29	36	1227	154	382	19	268	13
30	24	1460	282	405	11	281	16
31	25	1478	266	410	21	284	20
32	45	1711	364	425	21	293	14
33	65	1975	302	444	19	304	13
34	110	2213	362	455	23	311	13
35	133	2423	455	466	24	321	14
36	211	2735	395	481	19	328	11
37	351	2929	389	485	19	332	11
38	872	3071	357	491	17	335	11
39	1218	3198	365	498	17	338	11
40	1098	3278	388	502	16	340	11
41	505	3342	386	504	15	343	11
42	92	3423	415	508	16	345	13

**Figure 1** Smoothed centiles for birth weight: boys.**Figure 2** Smoothed centiles for birth weight: girls.

regular menstrual cycles and was certain of her menstrual history. The gestation of each infant was also assessed postnatally using the new Ballard score,<sup>15</sup> which had been evaluated in our neonatal unit and found to be applicable for Chinese infants. Only infants whose calculated gestation agreed within two weeks with that assessed postnatally were included.

### Measurements

All measurements were collected prospectively. Birth weight was measured by midwives within an hour of birth using an electronic weighing scale (Detecto Scale Co, USA), which was accurate to 5 g and calibrated before each measurement. Weighing was carried out after the infant had been thoroughly dried and the umbilical cord cut. Within 24–48 hours of birth, the crown-heel length and head circumference of each infant was measured using a neonatometer (Holtain, Dyfed, Wales, UK) and an inelastic tape measure (Harpenden Anthropometric Tape; Holtain) respectively. Crown-heel length was

measured from the top of the head to the sole of the foot with the baby lying supine. The head of the infant was held in the Frankfurt horizontal position with the lower edge of the bony orbit and the ear positioned in the same vertical plane. The hips and knees were extended using gentle force. The head circumference was the maximum circumference around the head at the level of the point just above the glabella anteriorly and the top of the occipital bone posteriorly. Three measurements were obtained and the largest one was recorded.

The mean, standard deviation, and 3rd, 5th, 10th, 25th, 50th, 75th, 90th, 95th, and 97th centiles of each variable at each gestation were computed separately for the male and female infants. Centile charts were constructed. The results were compared with those obtained for the 1982–1986 cohort.

### Statistical analysis

The LMS method using maximum penalised likelihood<sup>16</sup> was used to perform model fitting of the anthropometric centiles for body weight, crown-heel length, and head circumference.

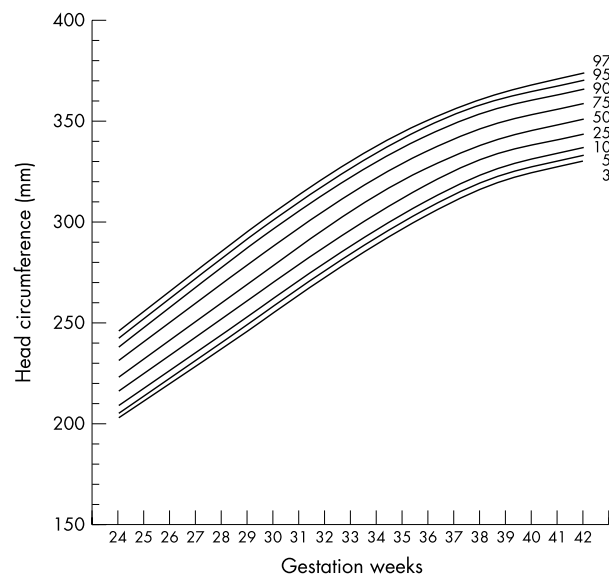


Figure 3 Smoothed centiles for head circumference: boys.

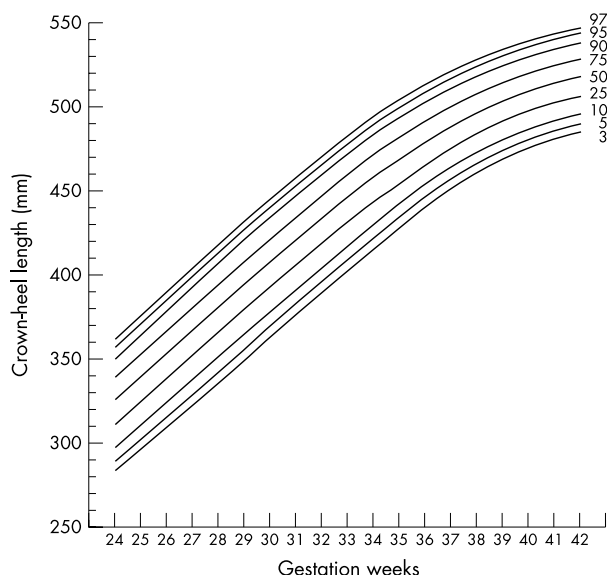


Figure 5 Smoothed centiles for crown-heel length: boys.

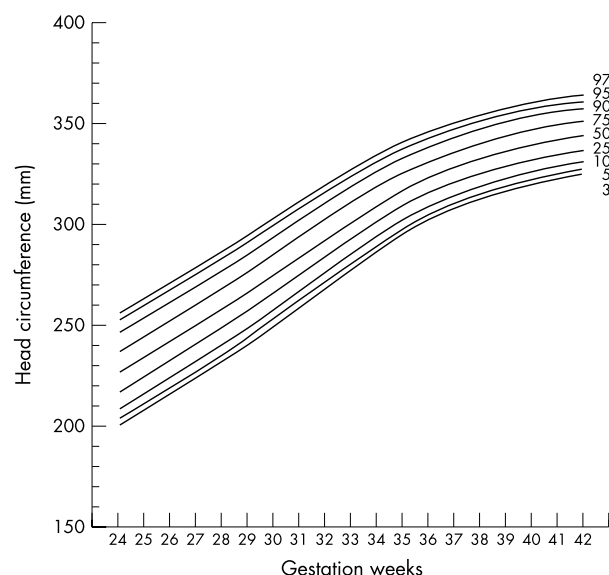


Figure 4 Smoothed centiles for head circumference: girls.

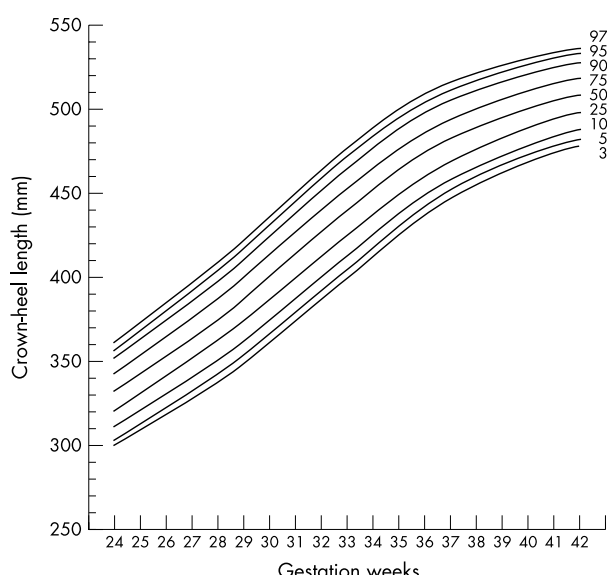


Figure 6 Smoothed centiles for crown-heel length: girls.

The LMS method estimates the measurement centiles in terms of three age-sex specific cubic spline curves: the L curve (Box-Cox power to transform the data that follow a normal distribution), M curve (median), and S curve (coefficient of variation). In brief, if  $Y(t)$  denotes an independent positive datum—for example, birth weight—at  $t$  gestation weeks, the distribution of  $Y(t)$  can be summarised by a normally distributed SD score ( $Z$ ) as follows:

$$Z = \frac{[Y(t)/M(t)]^{L(t)} - 1}{L(t)S(t)}$$

Once the  $L(t)$ ,  $M(t)$ , and  $S(t)$  have been estimated for each gestation  $t$ , the  $100\alpha$ th centile at  $t$  gestation weeks could be derived from

$$C_{100\alpha}(t) = M(t)[1 + L(t)S(t)Z_\alpha]^{1/L(t)}$$

where  $Z_\alpha$  is the  $\alpha$  centile of the normal distribution (for example for the 97th centile,  $\alpha = 0.75$  and  $Z_\alpha = 1.88$ ).

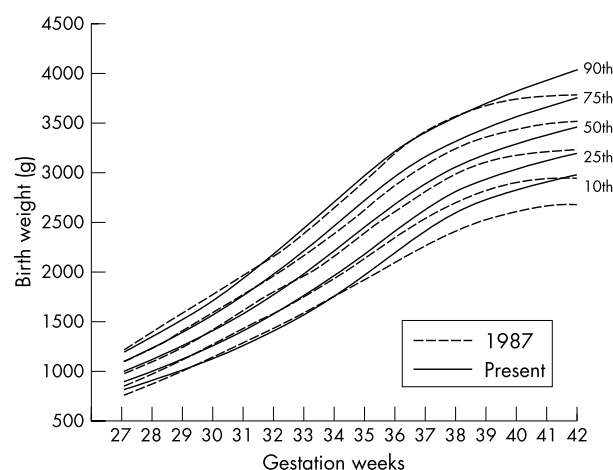
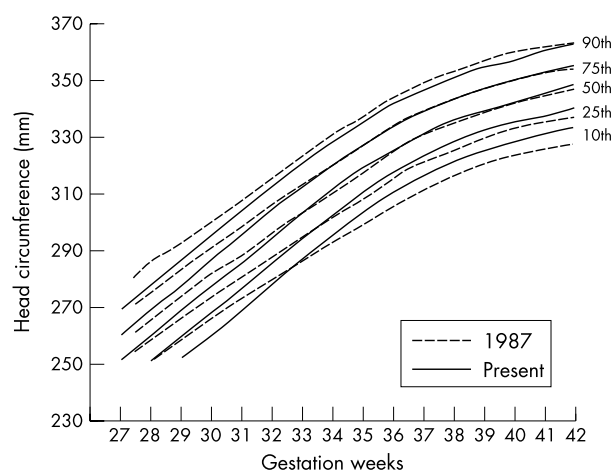
**RESULTS**

From October 1998 to September 2000, a total of 104 258 infants were born in Hong Kong: 76 230 were delivered in the public hospitals and 28 028 were delivered in the private hospitals. Measurements were obtained from a total of 10 032 singleton infants, or 9.6% of the total live born infants delivered during this period. Five infants were excluded because of discrepancy between the calculated gestation and that estimated postnatally. The remaining 10 027 infants included 9381 infants born in the public hospitals, and 646 infants born in the private hospitals. These represented 12.3% and 2.3% of the infants delivered in the public and private hospitals in Hong Kong respectively. After the addition of 307 infants with gestational age  $\leq 35$  weeks recruited in the subsequent nine months, and excluding infants with gestational age  $< 24$  weeks and  $> 42$  weeks, who were too few in number, a total of 10 339 infants (5478 boys, 4861 girls) were included for analysis. Gestation was estimated by early ( $< 20$  gestational

**Table 4** Comparisons of birth weight, crown-heel length and head circumference between the 1982–86 and present cohorts: combined data for boys and girls

Gestation (weeks)	n		Birth weight (g)			Crown-heel length (cm)			Head circumference (cm)		
	1982–86	Present	1982–86	Present	p Value	1982–86	Present	p Value	1982–86	Present	p Value
27	20	37	965 (148)	1024 (197)	0.247	36.6 (1.7)	36.4 (1.7)	0.673	25.6 (0.9)	25.1 (1.2)	0.109
28	30	46	1095 (157)	1140 (168)	0.245	37.4 (1.8)	37.7 (2.1)	0.522	26.4 (1.0)	26.0 (1.3)	0.157
29	45	62	1239 (210)	1259 (178)	0.596	38.5 (1.3)	39.1 (2.4)	0.131	26.9 (1.2)	26.9 (1.4)	1.000
30	41	75	1405 (197)	1471 (221)	0.113	39.7 (2.1)	40.7 (1.6)	0.005	27.8 (1.2)	28.1 (1.5)	0.273
31	52	66	1598 (208)	1578 (273)	0.663	41.1 (2.1)	41.9 (2.3)	0.054	28.5 (1.2)	28.7 (1.8)	0.492
32	88	81	1798 (196)	1793 (389)	0.915	42.6 (2.0)	42.9 (2.4)	0.377	29.3 (1.7)	29.6 (1.7)	0.253
33	97	154	1975 (217)	2022 (347)	0.233	43.8 (2.1)	44.6 (2.4)	0.007	30.0 (1.6)	30.6 (1.5)	0.003
34	125	211	2128 (247)	2223 (357)	0.009	45.5 (2.4)	45.5 (2.2)	1.000	30.7 (1.5)	31.1 (1.3)	0.011
35	154	281	2368 (294)	2470 (436)	0.010	46.7 (2.3)	46.9 (2.3)	0.386	31.5 (1.4)	32.1 (1.4)	<0.001
36	174	486	2620 (350)	2773 (410)	<0.001	47.1 (1.9)	48.1 (2.0)	<0.001	32.5 (1.3)	33.0 (1.3)	<0.001
37	416	783	2830 (376)	2997 (407)	<0.001	47.9 (2.5)	48.9 (1.9)	<0.001	33.2 (1.5)	33.4 (1.1)	0.010
38	938	1926	3025 (342)	3144 (388)	<0.001	48.7 (1.7)	49.6 (1.7)	<0.001	33.4 (1.3)	33.8 (1.2)	<0.001
39	1926	2522	3112 (315)	3246 (377)	<0.001	49.0 (1.7)	50.1 (1.7)	<0.001	33.8 (1.5)	34.1 (1.1)	<0.001
40	2145	2277	3192 (322)	3349 (400)	<0.001	49.4 (1.8)	50.6 (1.7)	<0.001	34.1 (1.5)	34.4 (1.2)	<0.001
41	1442	1049	3227 (331)	3433 (418)	<0.001	49.8 (1.7)	50.9 (1.7)	<0.001	34.3 (1.1)	34.7 (1.2)	<0.001
42	752	207	3265 (341)	3477 (417)	<0.001	49.8 (1.9)	51.3 (1.8)	<0.001	34.5 (1.1)	34.8 (1.2)	<0.001

Values are mean (SD).

**Figure 7** Comparison of smoothed centiles for birth weight between the 1982–1986 and present cohorts: combined data for boys and girls.**Figure 8** Comparison of smoothed centiles for head circumference between the 1982–1986 and present cohorts: combined data for boys and girls.

weeks) antenatal ultrasonography in 4497 (43.4%), and maternal last menstrual date (with confirmation by postnatal Ballard score) in 5852 (56.6%). Most (83%) of the infants were either first (44.2%) or second (38.8%) born infants. Para 3 infants constituted 11.9%, and only a small number (3.8%) were para 4 or above. This distribution of parity was similar to the general newborn population in the region. Table 1 summarises the characteristics of the parents.

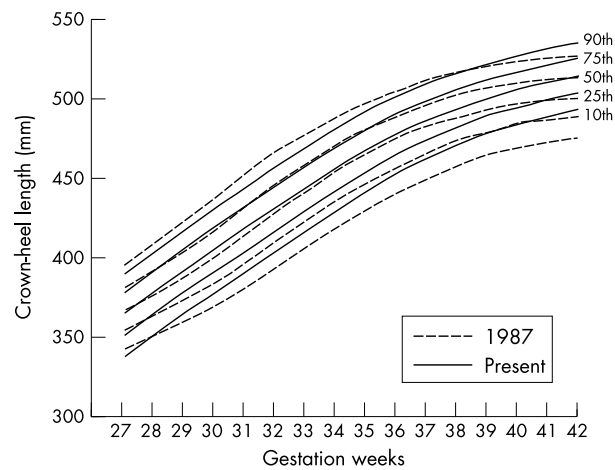
The interobserver agreement of the measurements between the two teams of investigators was assessed on the first 100 infants enrolled in the study, using the Bland-Altman procedure.<sup>17</sup> The results indicate that the agreement was acceptable, with only small discrepancies in the measurements of both crown-heel length (bias,  $-0.19$  mm; limit of agreement,  $-7.72$  to  $7.73$  mm) and head circumference (bias,  $1.37$  mm; limit of agreement,  $-4.40$  to  $7.13$  mm). The birth weight of each infant was measured twice by the midwives using the same electronic weighing scale. The two measurements showed perfect agreement (bias = 0).

Tables 2 and 3 show the distribution of gestational age and summary statistics (mean (SD)) for birth weight, crown-heel length, and head circumference of the boys and girls respectively. At each gestation, each of these measurements was normally distributed. At gestations of 36 weeks or more,

the boys consistently exceeded the girls in all three variables. Figures 1–6 show the gestation specific smoothed centile curve for each of the variables for each sex.

The combined data for boys and girls were compared with those for the 1982–1986 cohort, which were not broken down by sex (table 4, figs 7–9).<sup>13</sup> The present cohort of infants had a significantly higher mean birth weight at each gestation since 34 weeks. The mean differences ranged from 95 g (4.5%) at 34 weeks to 212 g (6.5%) at 42 weeks. This was accompanied by an upward shift in both the 10th and 90th centiles. No significant differences were observed between the two cohorts for the more preterm infants < 34 weeks gestation. The differences between the two cohorts were less obvious for crown-heel length and head circumference. Crown-heel length showed an upward shift of almost the entire 10th centile line from 28 to 42 weeks gestation, but the other centile lines showed such a shift only after 38 weeks gestation. Although the mean crown-heel lengths of the present cohort were significantly greater than those of the 1982–1986 cohort after 34 weeks gestation, the differences were small, ranging from 0.2 to 1.5 cm (0.4–3%). The head circumferences showed an upward shift of the 10th centile line after 33 weeks gestation, but the other centile lines were almost identical with those of the 1982–1986 cohort. Similarly to the crown-heel





**Figure 9** Comparison of smoothed centiles for crown-heel length between the 1982–1986 and present cohorts: combined data for boys and girls.

lengths, mean head circumferences in the present cohort were larger by a small, but significant, margin after 33 weeks gestation, with the differences ranging from 0.2 to 0.6 cm (0.6–1.9%).

## DISCUSSION

Accurate description of gestational age specific physical variables has three prerequisites: (a) reliable and accurate measurement, with not too many observers taking the measurements; (b) accurate recording of gestational age; (c) a sufficiently large sample of babies at various gestations to enable proper statistical description of the data. Most previous studies on gestation specific body size have used data documented in birth records. Using retrospective data has the advantage of enabling the inclusion of a large number of infants,<sup>18–20</sup> but often requires statistical exclusion of extreme outliers resulting from errors in documentation, erroneous estimation of gestation, or inaccurate measurements. These errors have led to appreciable bimodal distribution of birth weight at each gestation and the apparent excess of large preterm infants often observed in earlier studies.<sup>11–21–22</sup> Estimation of body length and head circumference is particularly prone to error because standard methods and precision instruments are not always used in the routine measurement of these variables in most maternity units. The prospective nature of our study allowed accurate determination of gestational age and measurement of the physical variables. As a result, there were no missing data, and exclusion of extreme outliers in data analysis was not required.

In general, for the construction of centile curves, the larger the sample size the greater the precision in the resulting centiles. We have calculated the minimum sample size at each gestation using the method described by Healy.<sup>23</sup> As in most similar studies, the number of infants at or near full term was well in excess of the minimum sample size, but the number of the very preterm infants, especially those below 28 weeks gestation, was small. On the basis of the weight distribution of the infants, our total sample size of 5478 would precisely yield the 95% confidence interval for the 97th centile within  $\pm 5\%$ , and  $\pm 6\%$  to  $\pm 14.2\%$  of its mean at each of 34th–42nd and 24th–33rd gestational week respectively for the male infants. For the female infants (total sample size 4861), the corresponding figures are  $\pm 6\%$ , and  $\pm 10\%$  to  $\pm 18\%$  respectively.<sup>24–25</sup>

Compared with the local data that we obtained in 1982–1986,<sup>13</sup> the present cohort had a significantly greater mean birth weight at each gestation since 34 weeks. This was accompanied by an upward shift of both the 10th and 90th centiles and much less terminal flattening of the curves after

term. The new chart thus redefines the cut off values for the diagnosis of both large and small for gestation. Therefore continuing to use the 1982–1986 standards will result in too few term or near term infants being classified as small for gestation, and too many classified as large for gestation. A similar secular trend in birth weights of term infants has been observed in other populations.<sup>20</sup> This trend was less obvious with crown-heel length and head circumference, which showed an upward shift only in the 10th centile line. Although significant, the differences in these two variables between the two cohorts are small and of doubtful clinical significance.

There are a number of controversies about the construction of gestation specific growth standards. The first is whether the data should be generated from a non-selected sample of the population or a selected sample of “healthy” subjects with no known factors affecting their growth. As one important application of these standards is to enable clinicians to identify subjects with growth problems, it has been pointed out by Cole<sup>26</sup> that it is illogical to construct a reference that is targeted at infants who are excluded by definition from the reference sample. It is also doubtful that a reference that truly represents “healthy” growth could ever be made available because many of the factors that affect fetal growth remain unidentified. Thus we did not exclude infants with antenatal factors affecting fetal growth, and the growth reference so constructed provides a neutral baseline for comparing groups without any assumption about the quality of the infants’ antenatal growth. The relative excess of infants delivered in public hospitals should not be a problem because infants born in public and private hospitals did not differ significantly in any of the three variables measured, making statistical weighting unnecessary. This is not surprising as the social infrastructure in Hong Kong allows easy and free access to health care for the entire population, and there is no distinct socioeconomic demarcation between the patients of public and private hospitals. In a previous study of Chinese newborns born in Mainland China, Taiwan, and the United States, Yip *et al*<sup>27</sup> showed that economic background had no significant influence on birth weight provided that the pregnant women could meet the basic health and nutrition requirements for adequate fetal growth.

Another controversy is whether preterm infants delivered by caesarean section should be excluded from data analysis. It has been argued that, as infants are delivered before term by interventional means often because of unsatisfactory intra-uterine growth or maternal or fetal complications that may compromise fetal growth, birth weight standards that include data obtained from these infants may be skewed towards the lighter end.<sup>28</sup> We compared the birth weights of our study infants delivered vaginally or by caesarean section before 37 weeks gestation, and did not observe in either sex any significant differences between the two groups at each gestational week. Exclusion of infants born by caesarian section was therefore unnecessary, and their data were included in the construction of the growth standards.

A third controversy is whether there should be one international standard for fetal growth rather than “local standards” representing the growth patterns of different populations.<sup>29–30</sup> Dunn<sup>30</sup> showed that the small size of infants born in developing countries is to a large extent the result of environmental factors such as maternal malnutrition, and suggested that their birth weights should be compared with an international perinatal growth reference which more truly reflects their growth potential than local charts. We compared the birth weights of our infants with those obtained from the non-indigenous population in Australia,<sup>19</sup> Norwegian infants,<sup>20</sup> and British infants in the Oxford area<sup>31</sup> (table 5). At gestation > 34 weeks, the 10th, 50th, and 90th centiles of both the Australian and Norwegian infants of either sex were substantially higher than those of our infants. At term ( $\geq 37$  weeks), the 50th centiles of our infants were less than those of the Norwegian infants by 280–306 g and the Australian infants

**Table 5** Differences of the 10th, 50th, and 90th centiles of gestation specific birth weights (g) of Chinese newborns from those of Australian,<sup>19</sup> Norwegian,<sup>20</sup> and British newborns<sup>31</sup>

Gestation (weeks)	Australian			Norwegian			British		
	10th centile	50th centile	90th centile	10th centile	50th centile	90th centile	10th centile	50th centile	90th centile
<b>Male</b>									
24	-32	28	95	13	73	120	48	88	205
25	-28	15	71	7	65	106	-18	20	41
26	-24	30	74	1	60	99	-44	-20	-6
27	-102	-14	70	13	71	105	-52	-34	-30
28	-93	29	70	17	89	130	-33	-21	-30
29	-100	61	96	40	121	166	-10	21	16
30	-88	58	122	77	158	197	32	58	52
31	10	67	113	120	207	243	60	97	83
32	-49	95	418	166	255	283	91	125	98
33	25	114	185	195	284	305	105	134	105
34	37	117	133	217	307	323	107	147	123
35	97	116	130	202	312	345	87	136	120
36	83	107	116	173	307	366	43	117	116
37	100	151	213	170	306	398	0	101	163
38	151	204	224	216	324	369	-29	94	164
39	173	217	241	253	352	391	-17	107	181
40	195	245	276	255	370	421	-5	115	176
41	214	277	310	239	357	415	4	97	130
42	161	235	289	186	305	364	-69	5	19
<b>Female</b>									
24	-13	-27	-46	-33	-17	-1	47	53	144
25	5	-18	-25	-35	-3	20	15	22	25
26	0	12	10	-30	12	45	20	17	20
27	-23	-2	29	-18	33	79	37	28	29
28	-80	2	45	10	72	130	70	52	45
29	-49	47	48	56	122	193	101	97	98
30	-9	62	75	106	172	250	146	132	135
31	-43	58	130	152	218	295	177	168	170
32	5	52	121	190	252	326	205	192	191
33	8	97	146	218	272	336	208	197	201
34	44	85	168	229	280	323	154	155	158
35	92	76	188	207	276	323	92	116	128
36	54	78	148	164	273	343	34	88	108
37	54	120	175	154	280	365	-16	70	115
38	110	168	232	180	303	382	-40	68	132
39	139	190	233	214	325	393	-31	80	153
40	174	220	264	239	345	404	-16	100	164
41	200	249	302	225	344	402	10	109	142
42	163	226	299	173	306	374	-47	46	69

by 120–230 g. The British data, which were for infants born 18–24 years ago, were more similar to our own, but there remained a significant difference of about 100 g between the 50th centiles of the two populations from 31 to 41 weeks gestation. A similar difference was also present between the 90th centile lines. We have previously compared the birth weights of our 1982–1986 cohort<sup>13</sup> with those obtained from a British<sup>32</sup> and an Australian<sup>33</sup> population over 30 years ago, and showed that the white infants were heavier throughout the third trimester, especially at gestations > 37 weeks. In both the 1982–1986 and the present cohorts, no maternal or environmental factors such as maternal malnutrition that could adversely affect fetal growth could be identified. It thus appears that, compared with white infants, there is a genuine genetic predisposition that leads to the smaller size of our infants.

Like all similar studies, the cross sectional anthropometric data obtained in this study do not reflect the intrauterine growth of the fetuses and are unsuitable for use in the evaluation of fetal growth velocity. Longitudinal study of individual fetuses would be required. Our study, however, provides an updated reference for evaluation of the size of Southern Chinese newborns of 24–42 weeks gestation and the identification of infants at risk of developing complications associated with excessively small or large size. Given the important relation between body size at birth and the future health of newborns,<sup>34–40</sup> these charts should be useful in the care of newborns of ethnic Chinese origin.

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## REFERENCES

- 1 **Wilcox AJ**, Skjaerven R. Birth weight and perinatal mortality: the effect of gestational age. *Am J Public Health* 1992;**82**:378–82.
- 2 **Wilcox AJ**, Russell IT. Birthweight and perinatal mortality. II. On weight-specific mortality. *Int J Epidemiol* 1983;**12**:319–25.
- 3 **Susser M**, Marolla FA, Fleiss J. Birth weight, fetal age and perinatal mortality. *Am J Epidemiol* 1972;**96**:197–204.
- 4 **Koops BL**, Morgan LJ, Battaglia FC. Neonatal mortality risk in relation to birth weight and gestational age: update. *J Pediatr* 1982;**101**:969–77.
- 5 **Williams RL**, Creasy RK, Cunningham GC, et al. Fetal growth and perinatal viability in California. *Obstet Gynecol* 1982;**59**:624–32.
- 6 **Allen MC**. Developmental outcome and followup of the small for gestational age infant. *Semin Perinatol* 1984;**8**:123–56.
- 7 **Commy JO**, Fitzhardinge PM. Handicap in the preterm small-for-gestational age infant. *J Pediatr* 1979;**94**:779–86.
- 8 **Melve KK**, Gjessing HK, Skjaerven R, et al. Infants' length at birth: an independent effect on perinatal mortality. *Acta Obstet Gynecol Scand* 2000;**79**:459–64.
- 9 **Commy JO**, Fitzhardinge PM. Handicap in the preterm small-for-gestational-age infant. *J Pediatr* 1979;**94**:779.
- 10 **Hong Kong Government Information Centre**. *Hong Kong Year Book 1999*. Hong Kong: Hong Kong Government Press, 2000.
- 11 **Lubchenco LO**, Hansman C, Dressler M, et al. Intrauterine growth estimated from liveborn birth-weight data at 24 to 42 weeks of gestation. *Pediatrics* 1963;**32**:793–80.
- 12 **Lubchenco LO**, Hansman C, Boyd E. Intrauterine growth in length and head circumference as estimated from livebirths at gestation ages from 26 to 42 weeks. *Pediatrics* 1966;**37**:403–8.
- 13 **Fok TF**, Lam TK, Lee N, et al. A prospective study on the intrauterine growth of Hong Kong Chinese babies. *Biol Neonate* 1987;**51**:312–23.
- 14 **Hong Kong Government**. *1996 Population By-Census Main Report*. Hong Kong: Hong Kong Government Press, 1997.
- 15 **Ballard JL**, Khoury JC, Wedig K, et al. New Ballard Score, expanded to include extremely premature infants. *J Pediatr* 1991;**119**:417.
- 16 **Cole TJ**, Freeman JV, Preece MA. British 1990 growth reference centiles for weight, height, body mass index and head circumference fitted by maximum penalized likelihood. *Stat Med* 1998;**17**:407–29.
- 17 **Bland JM**, Altman DG. Statistical methods for assessing agreement between two methods of clinical measurement. *Lancet* 1986;**1**:307–10.
- 18 **Arbuckle TE**, Wilkins R, Sherman GJ. Birth weight percentiles by gestational age in Canada. *Obstet Gynecol* 1993;**81**:39–48.
- 19 **Roberts CL**, Lancaster PA. Australian national birthweight percentiles by gestational age. *Med J Aust* 1999;**170**:114–18.
- 20 **Skjaerven R**, Gjessing HK, Bakkevig LS. Birthweight by gestational age in Norway. *Acta Obstet Gynecol Scand* 2000;**79**:440–9.
- 21 **Milner RD**, Richards B. An analysis of birth weight by gestational age of infants born in England and Wales, 1967 to 1971. *J Obstet Gynaecol Br Commonw* 1974;**81**:956–67.
- 22 **Gruenewald P**. Growth of the human fetus. I. Normal growth and its variation. *Am J Obstet Gynecol* 1966;**94**:1112–19.
- 23 **Healy HJ**. Statistics of growth standards. In: Falker F, Tanner JM, eds. *Human growth. A comprehensive treatise: methodology, ecological, genetic, and nutritional effects on growth*. New York: Plenum Press, 1986;**3**:47–58.
- 24 **Altman DG**, Chitty LS. Charts of fetal size. 1. Methodology. *Br J Obstet Gynaecol* 1994;**101**:29–34.
- 25 **Royston P**. Constructing time-specific reference ranges. *Stat Med* 1991;**10**:675–90.
- 26 **Cole TJ**. Some questions about how growth standards are used. *Horm Res* 1996;**45**(suppl 2):18–23.
- 27 **Yip R**, Li Z, Chong WH. Race and birth weight: the Chinese example. *Pediatrics* 1991;**87**:688–93.
- 28 **Lucas A**, Cole TJ, Gandy GM. Birthweight centiles in preterm infants reappraised. *Early Hum Dev* 1986;**13**:313–22.
- 29 **Wharton BA**. Sorrento studies of birthweight. Case for international reference data. *Acta Paediatr Scand Suppl* 1985;**319**:170–9.
- 30 **Dunn PM**. A perinatal growth chart for international reference. *Acta Paediatr Scand Suppl* 1985;**319**:180–7.
- 31 **Yudkin PL**, Aboualta M, Eyre JA, et al. New birthweight and head circumference centiles for gestational ages 24 to 42 weeks. *Early Hum Dev* 1987;**15**:45–52.
- 32 **Thomson AM**, Billewicz WZ, Hytten FE. The assessment of fetal growth. *J Obstet Gynaecol Br Commonw* 1968;**75**:903–16.
- 33 **Betheras FR**, White JG, Betheras GW. Intrauterine growth in an Australian population. *Aust N Z J Obstet Gynaecol* 1969;**9**:3–61.
- 34 **Barker DJ**, Martyn CN, Osmond C, et al. Growth in utero and serum cholesterol concentrations in adult life [see comments]. *BMJ* 1993;**307**:1524–7.
- 35 **Barker DJ**, Osmond C, Simmonds SJ, et al. The relation of small head circumference and thinness at birth to death from cardiovascular disease in adult life. *BMJ* 1993;**306**:422–6.
- 36 **Barker DJ**. Fetal growth and adult disease. *Br J Obstet Gynaecol* 1992;**99**:275–6.
- 37 **Barker DJ**, Godfrey KM, Fall C, et al. Relation of birth weight and childhood respiratory infection to adult lung function and death from chronic obstructive airways disease. *BMJ* 1991;**303**:671–5.
- 38 **Barker DJ**, Bull AR, Osmond C, et al. Fetal and placental size and risk of hypertension in adult life [see comments]. *BMJ* 1990;**301**:259–62.
- 39 **Barker DJ**, Winter PD, Osmond C, et al. Weight in infancy and death from ischaemic heart disease [see comments]. *Lancet* 1989;**2**:577–80.
- 40 **Barker DJ**, Osmond C, Golding J, et al. Growth in utero, blood pressure in childhood and adult life, and mortality from cardiovascular disease. *BMJ* 1989;**298**:564–7.



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